

# STATISTICAL THEORY. A CONCISE INTRODUCTION

## (Second Edition)

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### ERRATA

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- p.83, Corollary 5.1.  $g(\bar{Y}_n)$  is a consistent estimator of  $\theta$  (not  $g(\theta)$ ).
- p.93, l.-5, Sketch of the proof of Theorem 5.10. It should be “the sample mean of i.i.d. random variables  $(\ln f_\theta(Y_i))' \dots$ ” instead of  $\ln f_\theta(Y_i)$ .
- p.100, the first sentence in Section 5.11. Typo: the hypotheses should be  $H_0 : \theta = \theta_0$  vs.  $H_1 : \theta \neq \theta_0$ .
- p.120, Example 6.18. The null hypothesis should be rejected if  $T_n \left( \frac{\mu - y}{s/\sqrt{n}} \right) < \frac{1}{C+1}$  instead of  $\frac{C}{C+1}$ .
- p.120, Example 6.17. Table 6.1 should be placed here.
- p.121, Section 6.5.3. Typo: the hypotheses should be  $H_0 : \theta \in (\theta_0 - \epsilon, \theta_0 + \epsilon)$  versus  $H_1 : \theta \notin (\theta_0 - \epsilon, \theta_0 + \epsilon)$ .
- p.129. Question 7.1. The more accurate version is “Show that for any prior  $\pi(\theta)$  the Bayes risk of the Bayes rule  $\rho(\pi, \delta_\pi^*)$  is not larger than the minimax risk w.r.t. the same loss.”
- p.131, the beginning of the proof of Lemma 7.1. Typo: it should be “Let  $d = a_2^* - a_1^* \dots$ ” ( $d$  was missing).
- p.134, Example 7.19. Typo: a correct rejection region  $\Omega_{1C}(\mathbf{y})$  should be  $\Omega_{1C}(\mathbf{y}) = \{\mathbf{y} : \lambda(\mathbf{y}) = \frac{\mathbf{f}_{\theta_1}(\mathbf{y})}{\mathbf{f}_{\theta_0}(\mathbf{y})} \geq C\}$  (not  $\frac{\ln f_{\theta_1}(\mathbf{y})}{\ln f_{\theta_0}(\mathbf{y})} \geq C$ ).
- p.134, the fourth line in Section 7.7. Typo: it should be  $R(\theta, \delta_\pi^*)$  in the integrand instead of  $R(\theta_\pi^*, \delta)$ .
- p.178. The correct version is:  $RSS_{test}(h)/n_{test} - \sigma^2$  an unbiased estimator of  $AMSE(\hat{g}_h, g)$ , not  $RSS_{test}(h)/n_{test} + \sigma^2$ .
- p.213, solution of Exercise 6.1. The correct solution is:  
Calculate the sequential posterior distribution  $\pi^*(\theta|y_1, y_2)$  and compare it with its counterpart  $\pi(\theta|y_1, y_2)$  based on the entire data  $(y_1, y_2)$ . We have

$$\pi(\theta|y_1) = \frac{f(y_1|\theta)\pi(\theta)}{f(y_1)},$$

$$\pi^*(\theta|y_1, y_2) = \frac{f(y_2|\theta)\pi(\theta|y_1)}{f(y_2)} = \frac{f(y_2|\theta)f(y_1|\theta)\pi(\theta)/f(y_1)}{\int f(y_2|\theta)f_1(y_1|\theta)\pi(\theta)d\theta/f(y_1)} = \frac{f(y_2|\theta)f(y_1|\theta)\pi(\theta)}{\int f(y_2|\theta)f_1(y_1|\theta)\pi(\theta)d\theta} ,$$

whereas due to the independency of  $y_1$  and  $y_2$  given  $\theta$ ,

$$\pi(\theta|y_1, y_2) = \frac{f(y_1, y_2|\theta)\pi(\theta)}{f(y_1, y_2)} = \frac{f(y_1|\theta)f(y_2|\theta)\pi(\theta)}{\int f(y_1|\theta)f(y_2|\theta)\pi(\theta)d\theta} = \pi^*(\theta|y_1, y_2) .$$

- p.214, solution of Exercise 7.1, part 1. Typos: in the formula for Bayes risk there should be  $p(\text{Gamma}(\alpha, \beta), \hat{\lambda})$  and  $E_{\text{Gamma}(\alpha, \beta)}$  instead of  $p(\text{Beta}(\alpha, \beta), \hat{\lambda})$  and  $E_{\text{Beta}(\alpha, \beta)}$ .