# Poisson statistics for sequences modulo one 

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Based on joint works with C. Aistleitner (Graz), S. Baker (Birmingham) and N. Technau (Madison)

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## Outline

(1) Background

- Uniform Distribution Modulo One
- Gap Distribution
- Correlations
(2) Results
- The Sequences $\left(n^{\alpha}\right)$ and $\left(\alpha^{n}\right)$
- Proof Ideas


## Uniform Distribution Modulo One

- Let $\left(x_{n}\right)$ be a sequence of real numbers. We are interested in the distribution of the fractional parts $\left\{x_{n}\right\}$ in the unit interval $[0,1)$.
- We say that $\left(x_{n}\right)$ is uniformly distributed modulo 1 , if for every interval $I \subseteq[0,1)$, we have


In other words, every interval gets its proportional share of fractional parts.

- Example: Take $x_{n}=\sqrt{2} n$. The fractional parts are:
$0.414214,0.828427,0.242641,0.656854,0.0710678,0.485281$,
$0.899495,0.313708,0.727922,0.142136$,


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## Example - The Distribution of $(\sqrt{2} n) \bmod 1$

Figure: The first 100 fractional parts of the sequence $(\sqrt{2} n)$.


Figure: The distribution of the first 1000 fractional parts of the sequence $(\sqrt{2} n)$.

## The Weyl Criterion

## Theorem (The Weyl criterion, 1914)

The sequence $\left(x_{n}\right)$ is uniformly distributed mod 1 if and only if for every fixed integer $k \neq 0$, we have

$$
\lim _{N \rightarrow \infty} \frac{1}{N} \sum_{n=1}^{N} e^{2 \pi i k x_{n}}=0
$$

## Theorem (Bohl, Sierpiński, Weyl, 1909-10)

Let $\alpha \in \mathbb{R}$ be an irrational number. Then the sequence $(\alpha n)$ is uniformly distributed mod 1 .

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$$
\left|\frac{1}{N} \sum_{n=1}^{N} e^{2 \pi i k n \alpha}\right|=\frac{\left|1-e^{2 \pi i k N \alpha}\right|}{N\left|1-e^{2 \pi i k \alpha}\right|} \leq \frac{2}{N \mid 1-e^{2 \pi i k \alpha \mid}}=\frac{1}{N|\sin \pi k \alpha|}
$$

## Examples

## Theorem (Weyl, 1914)

Let $p(x)=a_{m} x^{m}+\cdots+a_{1} x+a_{0}$ be a polynomial with real coefficients, with at least one of the coefficients $a_{i}(i \geq 1)$ irrational. Then the sequence $(p(n))$ is u.d. mod 1.

- Example: Take $p(x)=\sqrt{2} x^{2}$. The sequence $\left(\sqrt{2} n^{2}\right)$ is u.d. $\bmod 1$.


## Theorem (Fejér, Csillag 1930) <br> For any $\alpha \neq 0$ and $\sigma>0$ with $\sigma$ not integer, the sequence $\left(\alpha n^{\sigma}\right)$ is u.d. $\bmod 1$ <br> - Example: The sequence $(\sqrt{n})$ is u.d. $\bmod 1$

## Theorem (Koksma, 1935)

The sequence $\left(\alpha^{n}\right)$ is u.d mod 1 for almost all $\alpha>1$.

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## Comparison of Uniformly Distributed Sequences



Figure: Some uniformly distributed sequences modulo one

- The definition of uniform distribution mod 1 does not capture the fact that some sequences look much more equidistributed than others. It also fails to fully capture the "randomness" of the sequences.


## Fine-Scale Statistics

- Uniform distribution modulo 1 measures the number of points in fixed intervals $I \subseteq[0,1)$.
- Subtler aspects of a sequence may be revealed by considering fine-scale statistics, i.e., statistics in intervals of length $1 / \mathrm{N}$ (the average gap), after restricting to the first $N$ elements of the sequence.
- To study "randomness" of a sequence, we can compare these statistics to those of random points in the unit interval.
- Let $X_{1}, X_{2}, \ldots, X_{N}$ be i.i.d. random variables uniformly distributed in the unit interval. The order statistics $X_{(1)} \leq X_{(2)} \leq \cdots \leq X_{(N)}$ are distributed like the first $N$ points of a Poisson point process $\mathcal{N}(t)$ on $\mathbb{R}$ (conditional on the event $\mathcal{N}(1)=N$ )
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- For the first $N$ elements of a sequence $\left(x_{n}\right)$, denote the ordered fractional parts by

$$
x_{(1)}^{N} \leq x_{(2)}^{N} \leq \cdots \leq x_{(N)}^{N} \leq x_{(N+1)}^{N}
$$

where by convention $x_{(N+1)}^{N}=1+x_{(1)}^{N}$.

- Let $\delta_{n}^{N}=N\left(x_{(n+1)}^{N}-x_{(n)}^{N}\right)$ be the normalized gaps.
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$$
\lim _{N \rightarrow \infty} \frac{1}{N} \#\left\{1 \leq n \leq N: \delta_{n}^{N} \in I\right\}=\int_{I} e^{-s} d s
$$

- Consider the gaps in the sequence $(\alpha n)$.


## Theorem (Sós, Surányi, Świerczkowski, 1957-1959)

Let $\alpha \in \mathbb{R}$. For any $N$, consider the gaps $\delta_{n}^{N}$ of the fractional parts of the sequence $(\alpha n)$. Then $\delta_{n}^{N}$ attains at most three distinct values.


Figure: The first 15 fractional parts of the sequence $\sqrt{2} n$. The (rescaled) gaps are: 1.06602, 1.06602, $0.441559,1.06602,1.06602,0.441559$, 1.06602, 1.06602, 1.50758, 1.06602, 1.06602, 0.441559, 1.06602, 1.06602, 1.50758.

- We say that an irrational $\alpha$ is Diophantine, if for all $\epsilon>0$ and all integers $p, q \neq 0$

$$
\left|\alpha-\frac{p}{q}\right|>\frac{C(\alpha, \epsilon)}{q^{2+\epsilon}} .
$$

By Roth's theorem all algebraic irrationals are Diophantine, and by Khinchin's theorem, almost all $\alpha \in \mathbb{R}$ are Diophantine.

- Conjecture (Rudnick-Sarnak, 1998): Let $d \geq 2$, and let $\alpha$ be a Diophantine irrational. Then the gap distribution of the fractional parts of the sequence $\left(\alpha n^{d}\right)$ is Poissonian.
- Rudnick-Sarnak's conjecture is open even for the simplest choices of $\alpha$, e.g., is the gap distribution of the fractional parts of $\left(\sqrt{2} n^{2}\right)$ Poissonian? We do not even have a "metric" result (i.e., a result holding for almost all $\alpha$ ).
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The gaps of $\left(\sqrt{2} n^{2}\right)$


Figure: The gap distribution of the first 10000 fractional parts of the sequence $\left(\sqrt{2} n^{2}\right)$.

- Conjecture: Let $\alpha>0$ such that $\alpha$ in not an integer, and $\alpha \neq 1 / 2$. Then the gap distribution of the fractional parts of the sequence $\left(n^{\alpha}\right)$ is Poissonian.
- Here again, not a single case is known. The case $\sigma=1 / 2$ is special:


## Theorem (Elkies-McMullen, 2004) <br> The gap distribution of the fractional parts of the sequence $(\sqrt{n})$ converges to a non Poissonian limit distribution.

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## The Sequence ( $n^{\alpha}$ )

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The gaps of $\left(n^{3 / 4}\right)$


Figure: The gap distribution of the first $10^{5}$ fractional parts of the sequence $\left(n^{3 / 4}\right)$.

The gaps of $(\sqrt{n})$


Figure: The gap distribution of the first $10^{7}$ fractional parts of the sequence $(\sqrt{n})$.

## Intervals around Points

- Given $N$ points $\left(x_{n}\right)$ in the unit interval and $s>0$, a random interval of the form $\left[x-\frac{s}{N}, x+\frac{s}{N}\right]$ is expected to contain $2 s$ points.
- When the centre of the interval is one of the given points $x_{m}$, it contains

$$
\#\left\{1 \leq n \leq N: n \neq m,\left\|x_{n}-x_{m}\right\| \leq \frac{s}{N}\right\}
$$

other points.


Figure: Counting the number of points in an interval around a point

- Averaging over all centres $x_{m}$, we get the pair correlation function

$$
R_{2, N}([-s, s])=\frac{1}{N} \#\left\{1 \leq n \neq m \leq N:\left\|x_{n}-x_{m}\right\| \leq \frac{s}{N}\right\}
$$

which measures the average number of (other) points in an interval around a point.

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$$
\lim _{N \rightarrow \infty} R_{2, N}([-s, s])=2 s
$$

## Smoothed Pair Correlation

- For $f \in C_{c}^{\infty}(\mathbb{R})$ we can equivalently work with the smooth pair correlation function

$$
R_{2, N}(f)=\frac{1}{N} \sum_{k \in \mathbb{Z}} \sum_{1 \leq n \neq m \leq N} f\left(N\left(x_{n}-x_{m}-k\right)\right)
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$$
\lim _{N \rightarrow \infty} R_{2, N}(f)=\int_{-\infty}^{\infty} f(x) d x
$$

## The Pair Correlation of $\left(\alpha n^{d}\right)$

## Theorem (Rudnick-Sarnak, 1998)

Let $d \geq 2$. For almost all $\alpha \in \mathbb{R}$, the sequence $\left(\alpha n^{d}\right)$ has Poissonian pair correlation.

- Very little is known for specific values of $\alpha$. For example, is the pair correlation of $\sqrt{2} n^{2}$ Poissonian?


Figure: The pair correlation measure of the first 2000 fractional parts of the sequence $\left(\sqrt{2} n^{2}\right)$.

## The Pair Correlation of $(\sqrt{n})$

## Theorem (El-Baz-Marklof-Vinogradov, 2013)

The pair correlation of the sequence $(\sqrt{n})_{\sqrt{n} \notin \mathbb{Z}}$ is Poissonian.


Figure: The pair correlation measure of the first 2000 fractional parts of the sequence $(\sqrt{n})_{\sqrt{n} \notin \mathbb{Z}}$.

- For $f \in C_{c}^{\infty}\left(\mathbb{R}^{k-1}\right)$, the $k$-level correlation is defined by

$$
R_{k, N}(f)=\frac{1}{N} \sum_{\mathbf{k} \in \mathbb{Z}^{k-1}} \sum_{\substack{1 \leq n_{1}, \ldots, n_{k} \leq N \\ \text { distinct }}} f(N(\Delta-\mathbf{k}))
$$

where

$$
\Delta=\left(x_{n_{2}}-x_{n_{1}}, x_{n_{3}}-x_{n_{2}}, \ldots, x_{n_{k}}-x_{n_{k-1}}\right) .
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- Poissonian $k$-level correlation:

- If all the $k$-level correlations are Poissonian, then the gap distribution is also Poissonian.
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## The Sequences $\left(n^{\alpha}\right)$ and $\left(\alpha^{n}\right)$

## Theorem (Technau-Y., 2020)

The sequence ( $n^{\alpha}$ ) has Poissonian $k$-level correlation for almost all $\alpha>4 k^{2}-4 k-1$.
> - This gives Poisson pair correlation for almost all $\alpha>7$. Can one improve the range of $\alpha$ all the way down to $\alpha>0$ ?

## Theorem (Aistleitner-Baker, 2020)

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## The Sequence $\left(\alpha^{n}\right)$

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> For almost all $\alpha>1$, all $k$-level correlations of $\left(\alpha^{n}\right)$ are Poissonian, and hence the gap distribution is Poissonian for almost all $\alpha>1$.

Theorem (Aistleitner-Baker-Technau-Y., 2020)
Let $\left(a_{n}\right)$ be a sequence of positive real numbers such that
 for all sufficiently large $n$. Then for almost all $\alpha>0$, all $k$-level correlations of ( $\left.e^{\alpha a_{n}}\right)$ are Poissonian, and hence the gap distribution is Poissonian for almost all $\alpha>0$

- The case $\left(\alpha^{n}\right)$ follows by setting $a_{n}=n$ and $\tilde{\alpha}=e^{\alpha}$. This cannot be used directly for the sequence ( $n^{\alpha}$ ) corresponding to $a_{n}=\log n$, since the condition $\lim _{n \rightarrow \infty} \frac{a_{n}}{\log n}=\infty$ fails to hold.


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Let $\left(a_{n}\right)$ be a sequence of positive real numbers such that $\lim _{n \rightarrow \infty} \frac{a_{n}}{\log n}=\infty$, and such that $a_{n+1}-a_{n} \geq n^{-C}$ for some $C>0$ for all sufficiently large $n$. Then for almost all $\alpha>0$, all $k$-level correlations of ( $\left.e^{\alpha a_{n}}\right)$ are Poissonian, and hence the gap distribution is Poissonian for almost all $\alpha>0$.

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## Theorem (Aistleitner-Baker-Technau-Y., 2020)

Let $\left(a_{n}\right)$ be a sequence of positive real numbers such that $\lim _{n \rightarrow \infty} \frac{a_{n}}{\log n}=\infty$, and such that $a_{n+1}-a_{n} \geq n^{-C}$ for some $C>0$ for all sufficiently large $n$. Then for almost all $\alpha>0$, all $k$-level correlations of ( $\left.e^{\alpha a_{n}}\right)$ are Poissonian, and hence the gap distribution is Poissonian for almost all $\alpha>0$.

- The case $\left(\alpha^{n}\right)$ follows by setting $a_{n}=n$ and $\tilde{\alpha}=e^{\alpha}$. This cannot be used directly for the sequence ( $n^{\alpha}$ ) corresponding to $a_{n}=\log n$, since the condition $\lim _{n \rightarrow \infty} \frac{a_{n}}{\log n}=\infty$ fails to hold.
- Consider the (smooth) pair correlation $R_{2, N}(f)$ of the sequence ( $e^{\alpha a_{n}}$ ), where $\alpha \in \mathcal{J}:=[A, A+1]$. Applying the Poisson summation formula and using the rapid decay of the Fourier coefficients, we get

$$
\begin{aligned}
R_{2, N}(f) & =\left(1-\frac{1}{N}\right) \hat{f}(0) \\
& +\frac{1}{N^{2}} \sum_{0 \neq|k| \leq N^{1+\epsilon}} \widehat{f}\left(\frac{k}{N}\right) \sum_{1 \leq n \neq m \leq N} e^{2 \pi i k\left(e^{\alpha a_{n}}-e^{\alpha a_{m}}\right)} \\
& +O\left(N^{-\infty}\right)
\end{aligned}
$$

and we would like to show that for almost all $\alpha>0$, the middle term vanishes as $N \rightarrow \infty$.

- By a standard Borel-Cantelli argument, it is enough to show that the variance of $R_{2, N}$ decays with a polynomial rate, which reduces to showing that

$$
\frac{1}{N^{4}} \sum_{\substack{0 \neq\left|k_{1}\right| \leq N^{1+\epsilon} \\ 0 \neq\left|k_{2}\right| \leq N^{1+\epsilon}}} \sum_{\substack{1 \leq m_{1} \neq m_{2} \leq N \\ 1 \leq n_{1} \neq n_{2} \leq N}}\left|\int_{\mathcal{J}} e^{2 \pi i \phi(\alpha)} d \alpha\right|=O\left(N^{-\delta}\right)
$$

for some $\delta>0$, where

$$
\phi(\alpha)=k_{1}\left(e^{\alpha a_{m_{1}}}-e^{\alpha a_{m_{2}}}\right)+k_{2}\left(e^{\alpha a_{n_{1}}}-e^{\alpha a_{n_{2}}}\right) .
$$

## A van der Corput type Lemma

## Lemma

Let $\phi: \mathcal{J} \rightarrow \mathbb{R}$ be a $C^{\infty}$-function. Fix $\ell \geq 1$, and suppose that $\phi^{(\ell)}(\alpha)$ has at most $C$ zeros, and that the inequality $\max _{1 \leq i \leq \ell}\left|\phi^{(i)}(\alpha)\right| \geq \lambda>0$ holds throughout the interval $\mathcal{J}$. If $\ell=1$, assume additionally that $\phi^{\prime}$ is monotone on $\mathcal{J}$. Then

$$
\int_{\mathcal{J}} e^{2 \pi i \phi(\alpha)} d \alpha<_{\ell, C} \lambda^{-1 / \ell}
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- The crux of the argument is therefore to prove that at each $\alpha \in \mathcal{J}$, at least some derivative of

is large ("repulsion principle").


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## Repulsion Principle

- To prove repulsion, note that

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- Assume for simplicity that $m_{1}>m_{2}>n_{1}>n_{2}$. Then if we denote the Vandermonde matrix by $M$, then

so that an appropriate lower bound for $\max _{1 \leq i \leq 4}\left|\phi^{(i)}(\alpha)\right|$ will follow from an upper bound for $\left\|M^{-1}\right\|_{\infty}$, which we will get from our assumptions.
- There are "degenerate" configurations, e.g., when $m_{1}=n_{1}$, where $\phi(\alpha)$ depends on fewer independent terms. These are handled combinatorially.


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