

# The metric theory of the pair correlation function of real-valued lacunary sequences

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**Abstract** Let  $\{a(x)\}_{x=1}^{\infty}$  be a positive, real-valued, lacunary sequence. This note shows that the pair correlation function of the fractional parts of the dilations  $\alpha a(x)$  is Poissonian for Lebesgue almost every  $\alpha \in \mathbb{R}$ . By using harmonic analysis, our result—irrespective of the choice of the real-valued sequence  $\{a(x)\}_{x=1}^{\infty}$ —can essentially be reduced to showing that the number of solutions to the Diophantine inequality

$$|n_1(a(x_1) - a(y_1)) - n_2(a(x_2) - a(y_2))| < 1$$

in integer six-tuples  $(n_1, n_2, x_1, x_2, y_1, y_2)$  located in the box  $[-N, N]^6$  with the “excluded diagonals”; that is,

$$x_1 \neq y_1, \quad x_2 \neq y_2, \quad (n_1, n_2) \neq (0, 0),$$

is at most  $N^{4-\delta}$  for some fixed  $\delta > 0$ , for all sufficiently large  $N$ .

## 1. Introduction

A sequence of points  $\{\theta_n\}_{n=1}^{\infty}$  is uniformly distributed modulo one if given any fixed interval  $I$  in the unit circle  $\mathbb{R}/\mathbb{Z}$ , the proportion of fractional parts  $\theta_n \bmod 1$  which lie in  $I$  tends to the length of the interval  $I$ ; that is,

$$\#\{n \leq N : \theta_n \bmod 1 \in I\} \sim \text{length}(I) \cdot N, \quad N \rightarrow \infty.$$

We study the *pair correlation function*  $R_2$ , defined for every fixed interval  $I \subset \mathbb{R}$  by the property that

$$\lim_{N \rightarrow \infty} \frac{1}{N} \#\left\{1 \leq m \neq n \leq N : |\theta_n - \theta_m| \in \frac{1}{N}I\right\} = \int_I R_2(x) dx,$$

assuming that the limit exists. For a random sequence of  $N$  elements—that is,  $N$  uniform independent random variables in  $[0, 1)$  (the Poisson model)—the limiting pair correlation function is almost surely  $R_2(x) \equiv 1$ .

There are very few positive results on the pair correlation function available for specific sequences, a notable exception being the fractional parts of  $\sqrt{n}$  [6]; a more

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tractable problem is to randomize (a “metric” theory, in the terminology of uniform distribution theory) by looking at random multiples<sup>1</sup>  $\theta_n = \alpha a(n) \bmod 1$ , for almost all  $\alpha$ . There is a well-developed metric theory of the pair correlation function for integer valued sequences  $\{a(n)\}_{n=1}^\infty$ , initiated in [11], where polynomial sequences such as  $a(n) = n^d$ ,  $d \geq 2$ , are studied, with several developments in the last few years; see, e.g., [2, 3, 5, 8–10, 12, 13]. In this note, we study the case of real-valued lacunary sequences: Let  $a(x) > 0$  be a lacunary sequence of positive reals—that is, there is some  $C > 1$  so that for all integers  $x \geq 1$ ,

$$a(x + 1) \geq Ca(x).$$

For instance, we can take  $a(x) = e^x$ . It is known that for almost all  $\alpha$ , the sequence  $\alpha a(x) \bmod 1$  is uniformly distributed mod one [7, Chapter 1, Corollary 4.3]. Here and throughout this note, “almost all” is meant with respect to the Lebesgue measure on  $\mathbb{R}$ .

**THEOREM 1.1**

*Assume that  $\{a(x)\}_{x=1}^\infty$  is a lacunary sequence of positive reals. Then the pair correlation function of the sequence  $\{\alpha a(x)\}_{x=1}^\infty$  is Poissonian for almost all  $\alpha$ .*

When  $a(x)$  takes integer values, it was shown in [12] that for almost all  $\alpha$ , the pair correlation function is Poissonian. The case of pair correlation of sequences of rationals  $x_n = a_n/b_n$  with  $a_n$  integer-valued and lacunary and  $b_n$  integer-valued and (roughly speaking) sufficiently small (e.g.,  $a_n/b_n = 2014^n / [\log \log n]$ ) was treated in [5]. Here we treat any real-valued sequences.

We will reduce the problem to giving a bound for the number of lattice points satisfying a Diophantine inequality: For  $M = N^{1+\varepsilon}$  and  $K = N^\varepsilon$ , let  $\mathfrak{S}(N)$  be the set of integer six-tuples with

$$1 \leq y_i \neq x_i \leq N, \quad 1 \leq |n_i| \leq M \quad (i = 1, 2),$$

satisfying

$$|n_1(a(x_1) - a(y_1)) - n_2(a(x_2) - a(y_2))| < K.$$

Assume that

$$(A) \quad \#\mathfrak{S}(N) \ll N^{4-\delta}.$$

**THEOREM 1.2**

*Let  $\{a(x)\}_{x=1}^\infty$  be a sequence of distinct positive reals. Assume that (A) holds for some  $\delta > 0$ . Then the pair correlation function of  $\alpha a(x)$  is Poissonian for Lebesgue almost all  $\alpha \in \mathbb{R}$ .*

1. A different notion of randomizing has recently been investigated in [1], which studies the pair correlation of the sequence  $\alpha^n \bmod 1$  with  $\alpha$  random.

In the case of integer-valued sequences, the almost sure convergence of the pair correlation function to the Poisson limit (metric Poisson pair correlation) follows [11, 12] from a similar bound for the equation

$$n_1(a(x_1) - a(y_1)) - n_2(a(x_2) - a(y_2)) = 0.$$

See [2, 4] for a streamlined criterion for metric Poisson pair correlation in terms of the additive energy of the sequence.

In Section 4, we verify that (A) holds for lacunary sequences; hence, we obtain Theorem 1.1.

### 2. The pair correlation function

To study the pair correlation function, we use a smooth count (cf. [11]): For  $f \in C_c^\infty(\mathbb{R})$  or  $f$  being an indicator function of a compact interval, set

$$F_N(x) = \sum_{j \in \mathbb{Z}} f(N(x + j)),$$

which is periodic and localized on scale  $1/N$ . For a sequence  $\{\theta_n\}_{n=1}^\infty \subset \mathbb{R}/\mathbb{Z}$ , we define its pair correlation function by

$$(2.1) \quad R_2(f, N)(\{\theta_n\}_{n=1}^\infty) = \frac{1}{N} \sum_{1 \leq m \neq n \leq N} F_N(\theta_n - \theta_m).$$

In particular, for a fixed sequence  $\{a_n\}_{n=1}^\infty$ , we take  $\theta_n = \alpha a_n \pmod 1$ , and abbreviate the pair correlation function  $R_2(f, N)(\{\theta_n\}_{n=1}^\infty)$ , having fixed  $f$ , by  $R_2(f, N)(\alpha) = R_2(\alpha)$ .

It suffices to restrict  $\alpha$  to lie in a fixed finite interval and to consider a smooth average: Let  $\rho \in C_c^\infty(\mathbb{R})$ ,  $\rho \geq 0$ , be a smooth, compactly supported, non-negative weight function, normalized to give a probability density:  $\int_{\mathbb{R}} \rho(\alpha) d\alpha = 1$ . We define a smooth average

$$(2.2) \quad \langle X \rangle = \int_{\mathbb{R}} X(\alpha) \rho(\alpha) d\alpha.$$

#### 2.1. The expected value

We verify later (see Lemma 2.2) that (A) implies that there exists  $\delta > 0$  such that for  $M = N^{1+\varepsilon}$  and  $K \ll N^\varepsilon$  we have that

$$(B) \quad \#\{1 \leq n \leq M, 1 \leq x \neq y \leq N : n|a(x) - a(y)| < K\} \ll N^{2-\delta}.$$

Assuming for now this statement is true, we know  $\langle R_2(f, N) \rangle$  asymptotically, as follows.

#### LEMMA 2.1

If (B) holds, then the expected value of  $R_2(f, N)(\alpha)$  is

$$\langle R_2(f, N) \rangle = \int_{-\infty}^\infty f(x) dx + O(N^{-\delta}).$$

*Proof*

Let  $f \in C_c^\infty(\mathbb{R})$ . By using Poisson summation, we have the expansion

$$F_N(x) = \sum_{j \in \mathbb{Z}} f(N(x + j)) = \frac{1}{N} \sum_{n \in \mathbb{Z}} \widehat{f}\left(\frac{n}{N}\right) e(n x)$$

with  $e(z) := e^{2\pi i z}$ , which gives

$$(2.3) \quad R_2(\alpha) = \frac{1}{N^2} \sum_{n \in \mathbb{Z}} \widehat{f}\left(\frac{n}{N}\right) S_{n,N}(\alpha),$$

where

$$S_{n,N}(\alpha) = \sum_{1 \leq x \neq y \leq N} e(\alpha n(a(x) - a(y))).$$

Therefore, the expected value is

$$\langle R_2 \rangle = \int_{-\infty}^{\infty} R_2(\alpha) \rho(\alpha) d\alpha = \frac{1}{N^2} \sum_{n \in \mathbb{Z}} \widehat{f}\left(\frac{n}{N}\right) \sum_{1 \leq x \neq y \leq N} \widehat{\rho}(n(a(x) - a(y))).$$

The zero mode  $n = 0$  gives a contribution of

$$\frac{1}{N^2} \widehat{f}(0) N(N - 1) = \int_{-\infty}^{\infty} f(x) dx (1 + O(1/N)).$$

We split the sum over non-zero modes into two terms: Those with  $1 \leq |n| \leq M = N^{1+\varepsilon}$ , and those with  $|n| > M$ . To treat the contribution of modes with  $|n| > M = N^{1+\varepsilon}$ , we use  $|\widehat{f}(x)| \ll x^{-A}$  and  $|\widehat{\rho}| \ll 1$  to bound that term by

$$\frac{1}{N^2} \sum_{|n| > M} \left(\frac{n}{N}\right)^{-A} \sum_{1 \leq x \neq y \leq N} 1 = \frac{N^A}{M^{A-1}} \ll \frac{1}{N^{1-\varepsilon}}$$

on choosing  $A = 2/\varepsilon$ .

To bound the contribution of modes with  $1 \leq |n| \leq M$ , we separate into a contribution of terms with  $|n(a(x) - a(y))| < K$  and the rest.

We use  $|\widehat{\rho}|, |\widehat{f}| \ll 1$  to obtain that the contribution of terms with  $|n(a(x) - a(y))| < K$  is

$$\ll \frac{1}{N^2} \#\{1 \leq n < N^{1+\varepsilon}, 1 \leq y \neq x \leq N : n(a(x) - a(y)) < N^\varepsilon\}.$$

By (B), this is  $\ll N^{-\delta}$ .

The contribution of terms with  $|n(a(x) - a(y))| > K$  is bounded using

$$|\widehat{\rho}(n(a(x) - a(y)))| \ll |n(a(x) - a(y))|^{-A} \leq K^{-A} = N^{-2}$$

and  $|\widehat{f}| \ll 1$  by

$$\frac{1}{N^2} \sum_{\substack{1 \leq |n| \leq M \\ 1 \leq x \neq y \leq N \\ |n(a(x) - a(y))| > K}} \left| \widehat{f}\left(\frac{n}{N}\right) \widehat{\rho}(n(a(x) - a(y))) \right| \ll \frac{1}{N^2} \sum_{\substack{1 \leq |n| \leq M \\ 1 \leq x \neq y \leq N}} \frac{1}{N^2} \leq \frac{M}{N^2},$$

which is  $\ll N^{-1+\varepsilon}$ . □

2.2. The condition (B)

LEMMA 2.2

If (A) holds for  $\delta > 0$  and  $K = N^\varepsilon$ , then (B) is satisfied with  $\delta' = \delta/2$  in place of  $\delta$  and  $K' = 2N^\varepsilon$  in place of  $K$ .

*Proof*

Let  $1_{[0, K']}$  be the indicator function of  $[0, K']$ , and let  $N \geq 1$  be large. With  $n, x, y$  and  $n_i, x_i, y_i$  constrained as in (B), we note that

$$\left( \sum_{n, x, y} 1_{[0, K']} (n |a(x) - a(y)|) \right)^2$$

equals

$$\begin{aligned} & \sum_{n_i, x_i, y_i} 1_{[0, K']} (|n_1(a(x_1) - a(y_1))|) 1_{[0, K']} (|n_2(a(x_2) - a(y_2))|) \\ & \leq \sum_{n_i, x_i, y_i} 1_{[0, K']} (|n_1(a(x_1) - a(y_1)) - n_2(a(x_2) - a(y_2))|) = \#S(N). \end{aligned}$$

Since  $\#S(N) \ll N^{4-\delta}$ , combining these considerations and taking a square root completes the proof. □

2.3. The variance

PROPOSITION 2.3

Assume that  $a(x)$  is a sequence of real numbers such that (A) holds. Then

$$\left\langle \left| R_2(f, N) - \int_{-\infty}^{\infty} f(x) dx \right|^2 \right\rangle \ll N^{-\delta}.$$

*Proof*

By Cauchy–Schwarz,

$$\begin{aligned} \left\langle \left| R_2(f, N) - \int_{-\infty}^{\infty} f(x) dx \right|^2 \right\rangle & \leq 2 \langle |R_2(f, N) - \langle R_2 \rangle|^2 \rangle \\ & \quad + 2 \left\langle \left| \langle R_2 \rangle - \int_{-\infty}^{\infty} f(x) dx \right|^2 \right\rangle. \end{aligned}$$

By Lemma 2.1,

$$\left\langle \left| \langle R_2 \rangle - \int_{-\infty}^{\infty} f(x) dx \right|^2 \right\rangle \ll N^{-2\delta}.$$

We now show that

$$(2.4) \quad \text{Var } R_2 = \langle |R_2(f, N) - \langle R_2 \rangle|^2 \rangle \ll N^{-\delta}$$

which will prove Proposition 2.3.

To prove (2.4), it suffices to show by (A) that

$$(2.5) \quad \text{Var } R_2 \ll_f \frac{\#\mathcal{S}(N)}{N^4}.$$

By using the expansion (2.3), the variance can be written as

$$(2.6) \quad \text{Var}(R_2) = \frac{1}{N^4} \sum_{(n_1, n_2) \in \mathbb{Z}^2 \setminus \{0\}} \widehat{f}\left(\frac{n_1}{N}\right) \widehat{f}\left(\frac{n_2}{N}\right) w(n_1, n_2, N),$$

where for integers  $n_1, n_2$ , we let

$$w(n_1, n_2, N) = \sum_{\substack{1 \leq x_1 \neq x_3 \leq N, \\ 1 \leq x_2 \neq x_4 \leq N}} \widehat{\rho}(n_1(a(x_3) - a(x_1)) - n_2(a(x_4) - a(x_2)))$$

and  $\rho$  as in (2.2).

Due to the rapid decay of  $\widehat{f}$ , the contribution from the range in which  $|n_1|$  or  $|n_2|$  exceeds  $M = N^{1+\varepsilon}$  is negligible, as we will argue now. We detail only the case  $\max\{|n_1|, |n_2|\} = n_1 \geq M$  since the other case can be done similarly. We observe the trivial bound  $|w(n_1, n_2, N)| \ll N^4$ . Moreover,

$$n_1 = n_1^{\varepsilon/2} n_1^{1-\varepsilon/2} \geq n_1^{\varepsilon/2} N^{1+\varepsilon/2-\varepsilon^2/2},$$

which, since  $\varepsilon$  is small, yields  $n_1 > n_1^{\varepsilon/2} N^{1+\varepsilon/3}$ . Hence, the contribution to the right-hand side of (2.6) arising from the terms with  $\max\{|n_1|, |n_2|\} = n_1 \geq M = N^{1+\varepsilon}$  and  $n_2 \neq 0$  is

$$\ll \frac{1}{N^4} \sum_{\substack{n_1, n_2 \in \mathbb{Z} \setminus \{0\} \\ |n_1| > N^{1+\varepsilon}}} \left(\frac{n_1}{N}\right)^{-18/\varepsilon} \sum_{n_2 \neq 0} \widehat{f}\left(\frac{n_2}{N}\right) N^4 \ll \frac{1}{N^4}.$$

Moreover, the terms satisfying  $\max\{|n_1|, |n_2|\} = n_1 \geq N^{1+\varepsilon}$ , and  $n_2 = 0$  are in absolute value

$$\ll \frac{1}{N^4} \sum_{|n_1| \geq N^{1+\varepsilon}} \widehat{f}\left(\frac{n_1}{N}\right) N^4 \ll \frac{1}{N^4}.$$

So, the upshot is that on the right-hand side of (2.6), the sum over all  $(n_1, n_2)$  with  $\max(|n_1|, |n_2|) > N^{1+\varepsilon}$  contributes  $\ll N^{-4}$ . By the rapid decay of  $\widehat{\rho}$ , we can dispose of the regime where  $|n_1(a(x_3) - a(x_1)) - n_2(a(x_4) - a(x_2))| \geq N^\varepsilon$ .

By bounding  $\widehat{\rho}$  trivially, we find that

$$\text{Var } R_2 \ll \frac{\#\mathcal{S}(N)}{N^4} + O\left(\frac{1}{N^4}\right).$$

Since  $\#\mathcal{S}(N) \geq N^3$ , we obtain (2.5). □

### 3. Almost everywhere convergence: Proof of Theorem 1.2

We now deduce almost everywhere convergence from a polynomial variance bound.

3.1. Preparations

We need a general property of the pair correlation function. Recall that for any sequence of points  $\{\theta_n\}_{n=1}^\infty \subset \mathbb{R}/\mathbb{Z}$ , we defined

$$R_2(f, N) = \frac{1}{N} \sum_{1 \leq j \neq k \leq N} F_N(\theta_j - \theta_k)$$

with  $F_N(x) = \sum_{j \in \mathbb{Z}} f(N(x - j))$ .

LEMMA 3.1

Suppose there is a strictly increasing sequence  $\{N_m\}_{m=1}^\infty \subseteq \mathbb{Z}_{\geq 1}$ , with

$$\lim_{m \rightarrow \infty} \frac{N_{m+1}}{N_m} = 1$$

so that for all  $f \in C_c^\infty(\mathbb{R})$ ,

$$(3.1) \quad \lim_{m \rightarrow \infty} R_2(f, N_m) = \int_{-\infty}^\infty f(x) dx.$$

Then we can pass from the sub-sequence to the set of all integers:

$$(3.2) \quad \lim_{N \rightarrow \infty} R_2(f, N) = \int_{-\infty}^\infty f(x) dx$$

for all  $f \in C_c^\infty(\mathbb{R})$ .

Proof

We will first deduce that (3.1) holds for the indicator functions

$$I_s(x) = \begin{cases} 1, & |x| < s/2, \\ 0, & \text{otherwise,} \end{cases}$$

by approximating with smooth functions, and show that (3.2) holds for the functions  $I_s$ , and then deduce by approximating a general even smooth  $f \in C_c^\infty(\mathbb{R})$  by linear combinations of  $I_s$  that (3.2) holds for all such  $f$ . Note that for odd smooth  $f \in C_c^\infty(\mathbb{R})$ , we have  $F_N(-x) = -F_N(x)$  which entails  $R_2(f, N) = 0$ , so the pair correlation function  $R_2(f, N)$  converges trivially to the right limit.

From the definition of  $R_2(I_s, N)$ , we have a monotonicity property: Let  $0 < \varepsilon < 1$ . If  $(1 - \varepsilon)N' < N < N'$  and  $N'' < N < (1 + \varepsilon)N''$ , then

$$(3.3) \quad (1 - \varepsilon)R_2(I_{(1-\varepsilon)s}, N'') \leq R_2(I_s, N) \leq \frac{1}{1 - \varepsilon}R_2(I_{s/(1-\varepsilon)}, N').$$

Indeed, using positivity of  $I_s$  (hence, of  $F_N$ ),

$$N \cdot R_2(I_s, N) = \sum_{1 \leq j \neq k \leq N} F_N(\theta_j - \theta_k) \leq \sum_{1 \leq j \neq k \leq N'} F_N(\theta_j - \theta_k).$$

Now if  $1 > N/N' \geq 1 - \varepsilon > 0$ , then since  $I_s$  is even and decreasing on  $[0, \infty)$ , we have

$$I_s(Ny) = I_s\left(N'y \cdot \frac{N}{N'}\right) \leq I_s(N'y(1 - \varepsilon)) = I_{s/(1-\varepsilon)}(N'y).$$

So

$$F_N(x) = \sum_{j \in \mathbb{Z}} I_s(N \cdot (x - j)) \leq \sum_{j \in \mathbb{Z}} I_{s/(1-\varepsilon)}(N' \cdot (x - j)) = \tilde{F}_{N'}(x),$$

where  $\tilde{F}_{N'}(y) = \sum_{j \in \mathbb{Z}} I_{s/(1-\varepsilon)}(N'(y - j))$ . Hence,

$$R_2(I_s, N) \leq \frac{N'}{N} R_2(I_{s/(1-\varepsilon)}, N') \leq \frac{1}{1-\varepsilon} R_2(I_{s/(1-\varepsilon)}, N'),$$

which proves the upper bound in (3.3). The lower bound of (3.3) follows from switching the roles of  $N$  and  $N''$  and inserting in the upper bound.

Next, fix  $\varepsilon \in (0, 1)$  small, let  $N \gg 1$ , and take  $m \gg 1$  so that

$$N_m < N_{m+1} < (1 + \varepsilon)N_m,$$

and so if  $N_m \leq N < N_{m+1}$ , then

$$(1 - \varepsilon)N_{m+1} < N < N_{m+1}, \quad N_m \leq N < (1 + \varepsilon)N_m.$$

Then for all  $s > 0$ , we have

$$(1 - \varepsilon)R_2(I_{(1-\varepsilon)s}, N_m) \leq R_2(I_s, N) \leq \frac{1}{1 - \varepsilon} R_2(I_{s/(1-\varepsilon)}, N_{m+1}).$$

Taking  $m \rightarrow \infty$ , we find by (3.1)

$$\limsup_{N \rightarrow \infty} R_2(I_s, N) \leq \frac{1}{1 - \varepsilon} \int_{-\infty}^{\infty} I_{s/(1-\varepsilon)} dx = \frac{s}{(1 - \varepsilon)^2}$$

and

$$\liminf_{N \rightarrow \infty} R_2(I_s, N) \geq (1 - \varepsilon) \int_{-\infty}^{\infty} I_{(1-\varepsilon)s} dx = (1 - \varepsilon)^2 s.$$

Since  $\varepsilon > 0$  is arbitrary, we finally obtain

$$\lim_{N \rightarrow \infty} R_2(I_s, N) = s = \int_{-\infty}^{\infty} I_s(x) dx$$

so that (3.2) holds for all indicator functions  $I_s$ . Therefore, (3.2) holds for all test functions  $f \in C_c^\infty(\mathbb{R})$ . □

### 3.2. Proof of Theorem 1.2

*Proof*

It suffices to show that for almost every  $\alpha$  in a fixed compact interval  $I$ , we have

$$(3.4) \quad \lim_{N \rightarrow \infty} R_2(f, N)(\alpha) = \int_{-\infty}^{\infty} f(x) dx$$

for all  $f \in C_c^\infty(\mathbb{R})$ .

Let  $\rho \in C_c^\infty(\mathbb{R})$  be a non-negative function majorizing the indicator function of the interval  $I$ :  $\mathbf{1}_I \leq \rho$ . Then from the variance bound of Proposition 2.3, we find that for some  $\delta > 0$ , for all  $f \in C_c^\infty(\mathbb{R})$ ,

$$\int_I \left| R_2(f, N_m)(\alpha) - \int_{-\infty}^{\infty} f(x) dx \right|^2 \rho(\alpha) d\alpha \ll_f N^{-\delta}.$$



Hence, for the sequence

$$N_m = \lfloor m^{2/\delta} \rfloor,$$

we have that for almost all  $\alpha \in I$ ,

$$(3.5) \quad \lim_{m \rightarrow \infty} R_2(f, N_m)(\alpha) = \int_{-\infty}^{\infty} f(x) dx$$

for all  $f$ . Indeed, for each fixed  $f$  set,

$$X_m(\alpha) = \left| R_2(f, N_m)(\alpha) - \int_{-\infty}^{\infty} f(x) dx \right|^2.$$

Then

$$\int_I X_m(\alpha) d\alpha \leq \int_{-\infty}^{\infty} X_m(\alpha) \rho(\alpha) d\alpha \ll \frac{1}{N_m^\delta} \ll \frac{1}{m^2}.$$

Therefore,

$$\int_I \left( \sum_{m \geq 1} X_m(\alpha) \right) d\alpha \leq \sum_{m \geq 1} \int_{-\infty}^{\infty} X_m(\alpha) d\alpha \ll_f \sum_{m \geq 1} \frac{1}{m^2} < \infty,$$

so that  $\sum_{m \geq 1} X_m(\alpha)$  converges for almost all  $\alpha \in I$ . Thus,

$$\lim_{m \rightarrow \infty} X_m(\alpha) = 0$$

for almost all  $\alpha$ —i.e., (3.5) holds for our specific  $f$  for almost all  $\alpha \in I$ .

By a diagonalization argument (see [11]) that involves selecting a countable and dense collection of functions  $f$ , there is a set of  $\alpha$  whose complement has measure zero so that (3.5) holds for all  $f$ . Since  $N_{m+1}/N_m \rightarrow 1$ , we can use Lemma 3.1 to deduce (3.4) holds, proving Theorem 1.2.  $\square$

### 4. Lacunary sequences

From now on, we assume that  $\{a(x)\}_{x=1}^\infty$  is a lacunary sequence of (strictly) positive reals—that is, there is some  $C > 1$  so that

$$a(x + 1) \geq Ca(x)$$

for all integers  $x \geq 1$ . Consequently, we have for all  $x \geq y \geq 1$  that

$$a(x) \geq C^{x-y}a(y).$$

We will show that (A) holds, hence proving that the pair correlation function of  $\{\alpha a(x) \bmod 1\}_{x=1}^\infty$  is Poissonian for almost all  $\alpha$  (that is, Theorem 1.1).

#### 4.1. The condition (A)

##### PROPOSITION 4.1

Denote by  $\mathcal{S}(N)$  the set of integer six-tuples with  $1 \leq y_i \neq x_i \leq N, 1 \leq |n_i| \leq M$  for  $i = 1, 2$  that satisfy

$$\left| n_1(a(x_1) - a(y_1)) - n_2(a(x_2) - a(y_2)) \right| < K.$$

Assume that  $\{a(x)\}_{x=1}^\infty$  is a lacunary sequence of positive real numbers. Further, suppose  $K = o(M)$  and  $N^\gamma \ll M \ll N^\Gamma$  for some  $0 < \gamma < \Gamma < 2$ . Then

$$\#\mathcal{S}(N) \ll MN^2(\log M)^2.$$

Note that Theorem 1.2 and the Proposition 4.1 together imply Theorem 1.1.

*Proof*

The proof is a modification of [12, Proposition 2]: We are given the inequality

$$(4.1) \quad |n_1(a(x_1) - a(y_1)) - n_2(a(x_2) - a(y_2))| < K.$$

We may assume that  $n_i > 0$ , and  $1 \leq y_i < x_i \leq N$ ,  $i = 1, 2$ , and that  $x_1 \geq x_2$ . In particular, we may then assume that  $x_1 \geq 4 \log_C M \gg \log N$  because the number of such tuples with  $x_1 \ll \log N$  is at most  $O(M^2(\log N)^4)$ , which is admissible (that is,  $o(MN^2(\log N)^2)$ ) if  $M = O(N^\Gamma)$  for  $\Gamma < 2$ .

We fix  $n_1, x_1, y_1$ , and first show that (recall  $x_1 \geq x_2$ )

$$(4.2) \quad x_1 - x_2 \leq 2 \log_C M.$$

Indeed, we have a lower bound,

$$n_1(a(x_1) - a(y_1)) \geq 1 \cdot (a(x_1) - a(x_1 - 1)) \geq \left(1 - \frac{1}{C}\right)a(x_1)$$

(since  $y \leq x_1 - 1$ ), and an upper bound,

$$n_2(a(x_2) - a(y_2)) \leq Ma(x_2) = a(x_1)M \frac{a(x_2)}{a(x_2 + (x_1 - x_2))} \leq \frac{M}{C^{x_1 - x_2}}a(x_1)$$

since  $a(x + h) \geq C^h a(x)$  for  $h \geq 1$ . Hence,

$$n_1(a(x_1) - a(y_1)) - n_2(a(x_2) - a(y_2)) \geq \left(1 - \frac{1}{C}\right)a(x_1) - \frac{M}{C^{x_1 - x_2}}a(x_1).$$

Assuming that  $x_1 - x_2 > 2 \log_C M$  gives in particular

$$1 - \frac{1}{C} - \frac{M}{C^{x_1 - x_2}} > 1 - \frac{1}{C} - \frac{1}{M} > \frac{1}{2} \left(1 - \frac{1}{C}\right) > 0$$

for sufficiently large  $N$ . The condition (4.1) now forces

$$\frac{1}{2} \left(1 - \frac{1}{C}\right) < \left(1 - \frac{1}{C}\right) - \frac{M}{C^{x_1 - x_2}} \leq \frac{K}{a(x_1)} \ll \frac{K}{C^{x_1}}$$

which forces  $x_1 \ll \log_C K \leq \varepsilon \log M$ , which we assumed was not the case. Thus, we may assume that  $x_1 - x_2 \leq 2 \log_C M$ , which forces  $x_2 \geq 2 \log_C M$  since  $x_1 > 4 \log_C M$ .

Now fix  $x_2$  as well; then  $n_2$  will be determined by  $y_2$  because

$$n_2 = \frac{n_1(a(x_1) - a(y_1))}{a(x_2) - a(y_2)} + O\left(\frac{K}{a(x_2) - a(y_2)}\right)$$

and since  $a(y)$  is lacunary,  $K/(a(x_2) - a(y_2)) = o(1)$  if  $y_2 \geq \log_C N$  because

$$a(x_2) - a(y_2) \geq a(x_2) - a(x_2 - 1) \geq a(x_2) \left(1 - \frac{1}{C}\right) \gg C^{x_2} \geq M^2$$

since  $x_2 > 2 \log_C M$ .

So we will be done if we show that there are  $\ll \log N$  many choices of  $y_2$  such that  $x_2 - y_2 > 2 \log_C M$ . Indeed, if there are two pairs  $(y_2, n_2)$  and  $(y'_2, n'_2)$  for which (4.1) holds (recall all other variables are now fixed), with  $x_2 - y_2 > 2 \log_C M$ ,  $x_2 - y'_2 > 2 \log_C M$ , then since

$$a(y_2) \leq \frac{a(x_2)}{C^{x_2 - y_2}} \leq \frac{a(x_2)}{M^2},$$

we find that (4.1) implies

$$\begin{aligned} n_1(a(x_1) - a(y_1)) &= n_2(a(x_2) - a(y_2)) + O(K) \\ &= n_2 a(x_2) \left( 1 + \frac{a(y_2)}{a(x_2)} + O\left(\frac{K}{n_2 a(x_2)}\right) \right) \\ &= n_2 a(x_2) \left( 1 + O\left(\frac{K}{M^2}\right) \right) \end{aligned}$$

since  $n_2 a(x_2) \geq a(x_2) \geq C^{x_2} \gg M^2$  if  $x_2 \geq 2 \log_C M$ , and  $a(x_2)/a(y_2) \geq C^{x_2 - y_2} \gg M^2$ . If  $n'_2, y'_2$  is another such pair, then we also find

$$n_1(a(x_1) - a(y_1)) = n'_2 a(x_2) \left( 1 + O\left(\frac{K}{M^2}\right) \right)$$

so that

$$n_2 a(x_2) \left( 1 + O\left(\frac{K}{M^2}\right) \right) = n'_2 a(x_2) \left( 1 + O\left(\frac{K}{M^2}\right) \right)$$

which gives

$$n'_2 = n_2 \left( 1 + O\left(\frac{K}{M^2}\right) \right) = n_2 + O\left(\frac{K}{M}\right)$$

since  $n_2 \leq M$ . Thus, for  $M \gg N^\nu$ , while  $K \ll N^\epsilon = o(M)$ , we obtain  $n'_2 = n_2$ . Substituting this information in (4.1), and recalling that  $n_1, x_1, y_1$  are fixed and  $n_2, x_2$  are determined now (up to  $\ll \log N$  many choices), we infer that there are  $\ll \log N$  many choices for  $y_2$ . This completes the proof.  $\square$

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